

Probability & Measures

A measure is an extension of the length, Area, or volume of subsets in 1-, 2-, or 3-dimensional Euclidean space. Given a sample space Ω , a measure is a set function defined on the subsets of Ω . Naturally, we define a measure μ on a σ -field as follows:

Definition: Let \mathcal{F} be a collection of subsets of a sample space Ω . \mathcal{F} is called a σ -field (σ -algebra) if and only if (iff) it has the following properties:

(i) The empty set $\emptyset \in \mathcal{F}$

(ii) If $A \in \mathcal{F}$, then $A^c \in \mathcal{F}$ [\mathcal{F} is closed under complementation]

(iii) If $A_i \in \mathcal{F}, i=1, 2, \dots$, then $\bigcup A_i \in \mathcal{F}$ [\mathcal{F} is closed under countable unions]

Note: If $A_1, A_2, \dots \in \mathcal{F}$, then $A_1^c, A_2^c, \dots \in \mathcal{F}$ and therefore

$\bigcup_{i=1}^{\infty} A_i^c \in \mathcal{F}$. Using DeMorgan's law, we have

$$\left(\bigcup_{i=1}^{\infty} A_i^c\right)^c = \bigcap_{i=1}^{\infty} A_i. \text{ That is, } \bigcap_{i=1}^{\infty} A_i \in \mathcal{F}.$$

Note: The collection of 2 sets $\{\emptyset, \Omega\}$ is a σ -algebra, called trivial σ -algebra.

Note: The important σ -algebra is the smallest one that contains all of the open sets in Ω .

Example 1: Consider Ω is finite or countable. Define
 $F = \{ \text{all subsets of } \Omega \text{ including } \Omega \}$

If Ω has n elements, then there are 2^n sets in F .
 For example, if $\Omega = \{1, 2, 3\}$, then F is a collection
 of $2^3 = 8$ sets:

$$\begin{array}{ccc} \{1\} & \{1, 2\} & \{1, 2, 3\} = \Omega \\ \{2\} & \{1, 3\} & \emptyset \\ \{3\} & \{2, 3\} & \end{array}$$

If Ω is uncountable, it is not an easy task to find F .

Example 2: $\Omega = (-\infty, \infty)$ - Real line. Then F can be
 chosen to include all sets of the form:

$$[a, b], (a, b], (a, b), [a, b)$$

Also from properties of F , F contains all sets taking
 unions and intersections of sets of the above forms.

Example 3: Let A be a non-empty subset of Ω ($A \subset \Omega$).
 Then $F = \{\emptyset, \Omega, A, A^c\}$ is a σ -algebra containing A
 (smallest σ -algebra containing A). This σ -algebra is denoted
 by $F = \sigma(\{A\}) = \sigma$ -algebra induced by A .

Def.: Let (Ω, \mathcal{F}) be a measurable space. A set function μ defined on \mathcal{F} is called a measure iff it has the following

(i) $0 \leq \mu(A) < \infty$ for any $A \in \mathcal{F}$

(ii) $\mu(\emptyset) = 0$

(iii) If $A_i \in \mathcal{F}$, $i = 1, 2, \dots$, and A_i 's are disjoint, then

$$\mu\left(\bigcup_{i=1}^{\infty} A_i\right) = \sum_{i=1}^{\infty} \mu(A_i)$$

The triple $(\Omega, \mathcal{F}, \mu)$ is called a measure space. If $\mu(\Omega) = 1$, then μ is called a probability space. It is denoted by P instead of μ . The triple (Ω, \mathcal{F}, P) is called a Probability space.

Example (Lebesgue measure): Consider the measure μ on $(\mathbb{R}, \mathcal{F})$ such that $\mu([a, b]) = b - a$ for every finite $[a, b]$. It is called Lebesgue measure. In this case, the integral of f is

$$\int f d\mu = \int \dots \int f(x_1, \dots, x_n) dx_1 \dots dx_n$$

$\mathcal{F} = \mathcal{B} = \text{Borel field.}$

Example (Counting measure): We may define the measure μ on (Ω, \mathcal{F}) as

$$\mu(A) = \# \text{ of elements in the subset } A. \text{ Here,}$$

$$\int f d\mu = \sum_{x \in \Omega} f(x).$$

Example (Probability measure): Let P be a Prob. measure on $(\mathbb{R}, \mathcal{F})$ defined by $F(x) = P((-\infty, x]) = P(X \leq x)$, $x \in \mathbb{R}$

Measurable functions

* $f: \Omega \rightarrow \mathbb{R}$, let $B \subset \mathbb{R}$. If $f^{-1}(B) = \{x \in \Omega: f(x) \in B\} \in \mathcal{F}$ for every Borel set B .

Then f is called a measurable function.

In probability theory, a measurable function is called a random element and denoted by X, Y, \dots

* $X \sim P_x =$ absolutely continuous with density p . Then $F_X(x) = P(X \leq x) = P((-\infty, x]) = \int_{-\infty}^x p(u) du$

$$\Rightarrow p(x) = F'_X(x)$$

* If X is a random variable with prob. measure P , then $EX = \int x dP = \int x p(x) dx$

Inferential Theory

Let $\Theta =$ parameter space $[\mathbb{R}^q - q\text{-dimensional vectors of real}]$

- $g(\theta) =$ quantity to be estimated, $g(\theta) \in \mathbb{R}$

- $\mathcal{P} = \{P_\theta: \theta \in \Theta\}$ - collection of Prob. dists.

- $X =$ observed data taking values in Ω with $P_\theta \in \mathcal{P}$

- $S: \Omega \rightarrow \mathbb{R}$ is an estimator

- $L: \Theta \times \mathcal{S} \rightarrow [0, \infty)$ loss function. It is denoted by $L(\theta, s)$ or $L(\theta, d)$ with $L(\theta, g(\theta)) = 0$

Examples: squared error loss (SEL): $L(\theta, d) = (g(\theta) - d)^2$

Absolute = = (AEL): $L(\theta, d) = |g(\theta) - d|$

Relative = = (REL): $L(\theta, d) = \left| \frac{d}{g(\theta)} - 1 \right|^p, p > 1$

For $K(\theta) > 0$, $L(\theta, d) = K(\theta) |g(\theta) - d|^p$ includes all others.

- Let $R(\theta, d) = E_{\theta} L(\theta, d(x))$ - Risk function. We say that estimator d_1 dominates d_2 if $R(\theta, d_1) \leq R(\theta, d_2)$ for all $\theta \in \Theta$ with strict inequality for some $\theta \in \Theta$.
- Estimators are chosen based on the following criteria:
 - Best within unbiased or invariant class of estimators
 - Best on the average (Bayes estimator)
i.e. minimizes $\int R(\theta, d) w(\theta) d\theta$

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 weighting average
 or prior dist.
 - Best under worst circumstances
i.e. minimizes $\text{Min}_{\theta \in \Theta} R(\theta, d)$

Convexity

Def.: A subset $A \in \mathbb{R}^n$ is convex if $\underline{x}, \underline{y} \in A \Rightarrow \delta \underline{x} + (1-\delta) \underline{y} \in A$
 for $0 < \delta < 1$ [A, B convex $\Rightarrow A \cap B$ convex].

Def.: Let A be convex and $f: A \rightarrow \mathbb{R}$. f is convex if
 $f(\delta \underline{x} + (1-\delta) \underline{y}) \leq \delta f(\underline{x}) + (1-\delta) f(\underline{y})$
 $\forall \underline{x}, \underline{y} \in A$ & $0 < \delta < 1$.

Theorem (Holder Inequality): For positive r.v.'s X and Y with finite means, $E(X^{\alpha} Y^{1-\alpha}) \leq (EX)^{\alpha} (EY)^{1-\alpha}$.

Theorem (Jensen's Inequality): Let $f: A \rightarrow \mathbb{R}$ be convex, then
 $f(EX) \leq E(f(X))$

Example: $f(x) = x^2$, convex $\Rightarrow EX^2 \geq (EX)^2 = \mu^2 \Rightarrow \sigma^2 \geq 0$.

Exponential family

Def.: A parametric family $\{P_\theta: \theta \in \Theta\}$ is called an exponential family iff $P(x|\eta) = P_\eta(x) = e^{\eta^T T(x) - A(\eta)} \cdot h(x)$ where $T \in \mathbb{R}^s$, with

$$\eta^T T = \sum_{i=1}^s \eta_i T_i \text{ and } A(\eta) = \log \int e^{\eta^T T} h(x) d\mu(x) < \infty$$

$$\tilde{\Omega} = \left\{ \eta \mid \int e^{\eta^T T} h(x) d\mu < \infty \right\} = \text{natural parameter space}$$

$$\eta = \begin{pmatrix} \eta_1 \\ \vdots \\ \eta_s \end{pmatrix}$$

$$\text{or } \tilde{\Omega} = \left\{ \eta \mid A(\eta) < \infty \right\}$$

Example: Let P_θ be the binomial dist. $B(n, \theta)$

$$P(x|\theta) = P_\theta(x) = e^{x \log \frac{\theta}{1-\theta} + n \log(1-\theta)} \binom{n}{x} \mathbb{I}_{\{0, 1, \dots, n\}}(x)$$

$$T(x) = x, \eta(\theta) = \log \frac{\theta}{1-\theta}, A(\theta) = -n \log(1-\theta) \text{ and}$$

$$h(x) = \binom{n}{x} \mathbb{I}_{\{0, 1, \dots, n\}}(x)$$

$$\Rightarrow P(x|\eta) = e^{\eta x - n \log(1+e^\eta)} \binom{n}{x} \mathbb{I}_{\{0, 1, \dots, n\}}(x) \quad \text{— Natural exp. family}$$

$$\Rightarrow \tilde{\Omega} = \mathbb{R}$$

Example: Suppose μ is Lebesgue measure on \mathbb{R} ; $s=1$, $T_1(x) = x$, $h = \mathbb{I}_{(0, \infty)}$

$$\text{Then } A(\eta) = \log \int_0^\infty e^{\eta x} dx = \begin{cases} \log(-\frac{1}{\eta}) & \text{if } \eta < 0 \\ \infty & \text{if } \eta \geq 0 \end{cases}$$

$$\Rightarrow P(x|\eta) = e^{\eta x - \log(-\frac{1}{\eta})} \mathbb{I}_{(0, \infty)}(x)$$

where $\eta \in \tilde{\Omega} = (-\infty, 0)$.

Example: Normal dist. $N(\mu, \sigma^2)$. Consider $\underline{\theta} = (\mu, \sigma^2)$

$$P(x|\underline{\theta}) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2\sigma^2}(x-\mu)^2}$$

$$= \frac{1}{\sqrt{2\pi}} e^{\frac{\mu}{\sigma^2}x - \frac{1}{2\sigma^2}x^2 - \left(\frac{\mu^2}{2\sigma^2} + \log \sigma\right)}$$

2-parameter exp. fam.

$$T_1 = x, T_2 = x^2, \eta_1(\underline{\theta}) = \frac{\mu}{\sigma^2}, \eta_2(\underline{\theta}) = -\frac{1}{2\sigma^2} \Rightarrow h(x) = \frac{1}{\sqrt{2\pi}}$$

$$\Rightarrow P(x|\underline{\eta}) = e^{\eta_1 x + \eta_2 x^2 - A(\eta_1, \eta_2)}, \quad A(\eta_1, \eta_2) = \frac{\eta_1^2}{4\eta_2} + \log(1/\sqrt{2\eta_2})$$

This leads to the fact that $P(x|\underline{\eta})$ is a natural exp. family of full rank with $\underline{\eta} = \mathbb{R} \times (0, \infty)$

Example (Multinomial family), $(X_0, X_1, \dots, X_k) \sim \text{Mult.}(n; p_0, \dots, p_k)$

$$f_{\underline{\theta}}(x_0, \dots, x_k) = \frac{n!}{x_0! x_1! \dots x_k!} p_0^{x_0} p_1^{x_1} \dots p_k^{x_k} \mathbb{I}_B(x_0, \dots, x_k)$$

$$\text{where } B = \{(x_0, x_1, \dots, x_k) : x_i \geq 0, \sum_{i=0}^k x_i = n\}$$

$$\text{Consider } \eta = (\log p_0, \log p_1, \dots, \log p_k), \quad h(x) = \frac{n!}{x_0! x_1! \dots x_k!} \mathbb{I}_B(x)$$

$$f_{\underline{\theta}}(x) = e^{\eta^T x} h(x), \quad x \in \mathbb{R}^{k+1}$$

\Rightarrow The above form does not give an exp. family of full rank, since there is no open set of \mathbb{R}^{k+1} contained in the natural parameter space. Reparameterization leads to Exp. family of full rank.

$$\text{Using } \sum_{i=0}^k x_i = n \text{ \& } \sum_{i=0}^k p_i = 1 \Rightarrow x_0 = n - \sum_{i=1}^k x_i, \quad p_0 = 1 - \sum_{i=1}^k p_i$$

$$\Rightarrow f_{\underline{\theta}}(x) = e^{\log p_0 (n - \sum_{i=1}^k x_i) + \sum_{i=1}^k \log p_i x_i} = e^{\sum_{i=1}^k \log \frac{p_i}{p_0} x_i + n \log p_0}$$

$$= e^{\sum_{i=1}^k \eta_i x_i - A(\underline{\eta})} \quad h(x), \quad x_{\text{sup}} = (x_1, \dots, x_k), \quad \underline{\eta} = \left(\log \frac{p_1}{p_0}, \dots, \log \frac{p_k}{p_0}\right)$$

$$A(\underline{\eta}) = -n \log p_0 \Rightarrow \text{Exp. family with full rank, } \underline{\eta} = \mathbb{R}^k$$

Moments of T_i 's

$$\text{Let } g(\underline{\eta}) = \int e^{\underline{\eta}' \underline{T}} h(x) d\mu = e^{A(\underline{\eta})}, \quad \underline{\eta} \in \mathbb{R}^s$$

Differentiating w.r.t η_j gives

$$e^{A(\underline{\eta})} \frac{\partial A(\underline{\eta})}{\partial \eta_j} = \int \frac{\partial}{\partial \eta_j} e^{\sum_{i=1}^s \eta_i T_i} h(x) d\mu$$

$$= \int T_j(x) e^{\sum_{i=1}^s \eta_i T_i} h(x) d\mu$$

$$\Rightarrow \frac{\partial A(\underline{\eta})}{\partial \eta_j} = \int T_j(x) e^{\sum_{i=1}^s \eta_i T_i - A(\underline{\eta})} d\mu$$

$$= \int T_j(x) p(x | \underline{\eta}) d\mu$$

$$\boxed{\frac{\partial A(\underline{\eta})}{\partial \eta_j} = E(T_j)}$$

$$\text{Now, } E T_i = \frac{\partial A(\underline{\eta})}{\partial \eta_i} = \int T_i e^{\underline{\eta}' \underline{T} - A(\underline{\eta})} h(x) d\mu$$

$$\frac{\partial^2 A(\underline{\eta})}{\partial \eta_j \partial \eta_i} = \int T_i e^{\underline{\eta}' \underline{T} - A(\underline{\eta})} \left(T_j - \frac{\partial A(\underline{\eta})}{\partial \eta_j} \right) h(x) d\mu$$

$$= E(T_i T_j) - E(T_i) E(T_j)$$

$$= \text{Cov}(T_i, T_j)$$

The moment generating function (MGF) of $\underline{T} = (T_1, \dots, T_s)$ is

$$M_{\underline{T}}(\underline{u}) = E e^{\underline{u}' \underline{T}} = \int e^{\underline{u}' \underline{T}} e^{\underline{\eta}' \underline{T} - A(\underline{\eta})} h(x) d\mu(x)$$

$$= \int e^{(\underline{u} + \underline{\eta})' \underline{T}} h(x) d\mu(x) / e^{A(\underline{\eta})}$$

$$= e^{A(\underline{u} + \underline{\eta})} / e^{A(\underline{\eta})} = e^{A(\underline{u} + \underline{\eta}) - A(\underline{\eta})}$$